Probability and Statistics Lecture 2 Random variables

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# Plan

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5 Expectation and variance

A **random variable** : quantity which depends only upon the result of the (random) experiment :

- number of "6" obtained when throwing 3 dices;
- number of phone calls during one hour ...

## Definition 1.1

Let  $(\Omega, \mathcal{T}, P)$  be a probability space. Let  $E \subset \mathbb{R}^d$  be given. Any map

$$X:\Omega o E, \omega \in \Omega o X(\omega)$$

is called a <u>random variable</u> on the probability space, valued in E.

If  $E \subset \mathbb{R}$ , we say that is a <u>real random variable</u>; If  $E \subset \mathbb{R}^d$ , we say that it is a <u>random vector</u>. So be careful : a random vector is a map.

# Distribution function of rv

## Definition 2.1

Distribution function of the real rv X :

$$F_X(x) = P(X \le x), \ \forall x \in I\!\!R$$

## **Proposition 2.1**

Properties : we have

**1** 
$$0 \le F_X(x) \le 1$$
,  $\lim_{x \to \infty} F_X = 1$ ,  $\lim_{x \to \infty} F_X = 0$ ;

•  $F_X$  is right continuous, that is  $F_X(x) = F_X(x^+)$  (right limit at x).

## **Proposition 2.2**

We have

$$P(a < X \le b) = F_X(b) - F_X(a)$$

#### **Proposition 2.3**

We have

$$P(X = x) = F_X(x) - F_X(x^-)$$

wher  $F_X(x^-)$  is teh left limit at x.

Important : if  $F_X$  is a continuous function, then the probability  $\overline{P(X = x)}$  is zero for any real x. In that case, that is if the distribution function is continuous, we have

 $P(a \leq X \leq b) = P(a < X \leq b) = P(a \leq X < b) = P(a < X < b)$ 

## **Definition 2.2**

Let X be a rv taking at most a countable number of values,  $X(\Omega) = \{x_1, x_2, ...\}$ . Then we say that X is a <u>discrete rv</u>.

### **Definition 2.3**

Let X taking an infinite non countable number of values. If  $F_X$  is continuous, we say that <u>X is a continuous rv</u>.

# Mass and density functions

## **Definition 3.1**

Let X be a discrete rv, 
$$X(\Omega) = \{x_1, x_2, ...\}$$
.  $p_X$ :

$$p_X(x_k) = P(X = x_k), \ k = 1, 2, ...$$

= mass function of X = probability law of X.

## **Definition 3.2**

Let X be a <u>continuous rv</u>. The function  $f_X$  defined by (if it exists)

$$f_X(x) = \frac{d}{dx} F_X(x)$$

is called the probability density function of X.

<u>Thus</u>  $f_X(x)dx \sim$  probability that rv X takes value x in a small interval of length dx around x.

# Properties : i) $f_X(x) \ge 0$ , because $F_X$ is increasing; ii) $\int_{-\infty}^{x} f_X(t) dt = F_X(x)$ $- > \int_{-\infty}^{+\infty} f_X(x) dx = F_X(+\infty) = 1$

A positive function with this property is called a density (function). Important property

$$P(a < X \le b) = F_X(b) - F_X(a) = \int_a^b f_X(x) dx$$

The probability that X belongs to the interval (a, b] is given by the area under the curve of  $f_X$  from a to b.

Random experiment E. Let  $A \subset \Omega$  and let X be the indicator rv of this event A, that is X = 1 if we get A and 0 otherwise. We say that X follows a Bernoulli law with parameter p, where p = P(A) is the probability of a success. We have

$$\begin{array}{cccc} x & 0 & 1 & \Sigma \\ p_X(x) & 1-p & p & 1 \end{array}$$

We may also write

$$p_X(x) = p^x q^{1-x}$$
,  $x = 0, 1$ 

where q = 1 - p. Here p is a known parameter.

We repeat the previous experiment E n times.

We then say that the trials  $E_1, ..., E_n$  form <u>Bernoulli trials</u> if

a) these trials are independent and

b) the probabillity of a success is the same for each trial.

## **Binomial law**

X : number of success with n Bernoulli trials.

X follows a bionomial law with parameters n and p, where p is the probability of one success. We have  $X(\Omega) = \{0, 1, ..., n\}$ . We write  $X \sim B(n, p)$ . Probability law of  $X \sim B(n, p)$  :

$$p_X(k) = C_k^n p^k q^{n-k}$$
, for  $k = 0, 1, ..., n$ 

# Geometric Law

X : number of necessary Bernoulli trials in order to obtain the first success. Then  $X(\Omega) = \{1, 2, ...\}$ . X follows a geometric law with parameter  $p : X \sim Geom(p)$ . In that case, we have

$$p_X(k) = q^{k-1}p$$
, pour  $k = 1, 2, ...$ 

We have :

$$F_X(n) = ... = 1 - q^n$$
 et  $P(X > n) = q^n$ ,  $n = 1, 2, ...$ 

As a last remark, it may happen that we can also define another X denoted by Y, to be the number of Bernoulli trials before getting the first success. In that case  $Y(\Omega) = \{0, 1, ..\}$  and function  $p_Y$  becomes

$$p_Y(k) = q^k p$$
,  $k = 0, 1, ...$ 

Then Y = X - 1.

## **Definition 4.1**

Let X be a drv with  $X(\Omega) = \{0, 1, ...;\}$  and

$$p_X(k) = \frac{e^{-\alpha} \alpha^k}{k!}$$
 for  $k = 0, 1, ...$ 

We then say that X follows a Poisson law with parameter  $\alpha > 0$ . We write  $X \sim Poi(\alpha)$ .

# CRV/ Uniform Law

We choose at random a number X in the interval [a, b]. If the density function of X is given by

$$f_x(x) = rac{1}{b-a}$$
,  $a \le x \le b$ 

we say that X follows a uniform law on the interval [a, b]. We write  $X \sim U(a, b)$ . Distribution function of X :

$$F_X(x) = 0$$
 if  $x < a$ ,  $\frac{x-a}{x-b}$  if  $a \le x \le b$ , 1 if  $x > b$ 

If  $[c, d] \subset [a, b]$ , then

$$P(c < X \le d) = \frac{d-c}{b-a}$$

The probability for X to be in a given sub interval only depends on the length of this interval.

We say that X defined on  $[0, +\infty)$  follows an exponential law with parameter  $\lambda$  if

$$f_X(x) = \lambda e^{-\lambda x}, x \ge 0$$

We write  $X \sim Exp(\lambda)$ . This is a density. The distribution function is given by

$$F_X(x) = e^{-\lambda x}$$
 if  $x \ge 0$  and 0 if  $x < 0$ 

In particuliar,

$$P(X > 0) = e^{-\lambda x}, x \ge 0$$

# CRV/ Gamma law

Gamma function  $\Gamma(.)$  :

$$\Gamma(\alpha) = \int_0^{+\infty} x^{\alpha-1} e^{-x} dx$$

pour  $\alpha > 0$ . I.B.P - >  $\Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1)$ , if  $\alpha > 1$ If  $\alpha = n = 2, 3, ...,$  then

$$\Gamma(n) = (n-1)!$$

Also :

$$\Gamma(1/2) = \sqrt{\pi}$$

Let X : a positive rv. If its density function is given by :

$$f_X(x) = rac{(\lambda x)^{lpha - 1} \lambda e^{-\lambda x}}{\Gamma(lpha)} ext{ for } x \geq 0,$$

we say that X follows a Gamma law wth parameters  $\alpha > 0$  and  $\lambda > 0$ . We write  $X \sim G(\alpha, \lambda)$ .

#### Remarks 5.1

i)  $\alpha$  : shape parameter;  $\lambda$  : scale parameter. As  $f_X$  has a shape which varies rapidly when parameter  $\alpha$  takes different values, Gamma law is very often used in applications. ii) Gamma law = exponential law when  $\alpha = 1$ . iii) If  $\alpha = n/2$ ,  $n \in \mathbb{N}$  and  $\lambda = 1/2$ , Gamma law is also called khi-square law with 2 to n degrees of freedom. Let X be a rv with  $X(\Omega) = \mathbb{R}$ . If the density function of X is of the form

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp\{-\frac{(x-\mu)^2}{2\sigma^2}\}, x \in \mathbb{R}$$

we say that X follows a gaussian or normal law with parameters  $\mu$  and  $\sigma^2$  with  $\sigma > 0$ . We write  $X \sim N(\mu, \sigma^2)$ . It is also called Laplace-Gauss law. Parameter  $\mu$  is a position paramter, while  $\sigma$  is a scale parameter. All gaussian laws have a bell shape. If  $\mu = 0$  and  $\sigma = 1$ : X follows a <u>standard normal law</u>. Its density function is given by

$$\phi(z) \equiv \frac{1}{\sqrt{2\pi}} exp(-z^2/2), z \in \mathbb{R}$$

and its distribution function denoted by  $\Phi$ , is

$$\Phi(z) = \int_{-\infty}^{z} \phi(y) dy = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{z} e^{-y^2/2} dy$$

IF  $X \sim N(\mu, \sigma^2)$ , -> distribution function in terms of  $\Phi$  :

$$F_X(x) = \Phi(\frac{x-\mu}{\sigma})$$

corresponding density function given by

$$f_X(x) = \frac{1}{\sigma}\phi(\frac{x-\mu}{\sigma})$$

That is we can obtain everything from the standard law N(0, 1).

## **Definition 6.1**

The expectation or the mean of a rv X, denoted by E(X) or equivalently by  $\langle X \rangle$  is defined by

$$E(X) \equiv \mu_X = \begin{cases} \sum_{k=1}^{\infty} x_k p_X(x_k) \text{ if } X \text{ is discrete} \\ \\ \int_{-\infty}^{\infty} x f_X(x) dx \text{ if } X \text{ is continuous.} \end{cases}$$

The expectation is linear.

## Proposition 6.1

Let X be a rv and Y = g(X). Then we have

$$E(Y) = \begin{cases} \sum_{k=0}^{\infty} g(x_k) p_X(x_k) \text{ if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} g(x) f_X(x) dx \text{ if } X \text{ is continuous} \end{cases}$$

#### Definition 6.2

The variance of rv X is defined by

$$Var(X) = \sigma_X^2 = E((X - E(X)^2))$$

If X has a density  $f_X$ , then we have

$$\sigma_X^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) dx$$

with  $\mu = E(X)$ . The standard deviation is defined by

$$STD(X) = \sigma_X = \sqrt{Var(X)}$$

We have

$$Var(aX + b) = a^2 Var(X)$$

and

$$Var(X) = E(X^2) - (E(X))^2$$

Law	Parameters	Mean	Variance
Bernoulli	р	р	pq
Binomial	n et p	np	npq
Geometric	р	1/p	$q/p^2$
Poisson	α	α	α

Law	Parameters	Mean	Variance
Uniform	[a, b]	(a+b)/2	$(b-a)^2/12$
Exponential	λ	$1/\lambda$	$1/\lambda^2$
Gaussian	$\mu$ and $\sigma^2$	μ	$\sigma^2$

## Theorem 6.1

Bienaymé-Tchebychev inequality Let X be a rv with mean  $\overline{E(X)} = \mu$  and variance  $Var(X) = \sigma^2$ . Then for any a > 0, we have

$$\mathsf{P}(|X-\mu| \ge \mathsf{a}) \le rac{\sigma^2}{\mathsf{a}^2}$$

#### Theorem 6.2

Markov inequality Let X be a positive rv. Then, for any a > 0

$$P(X \ge a) \le \frac{E(X)}{a}$$